

Doubly robust estimation of treatment effects under covariate dependent left truncation and right censoring

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Introduction

- Outcome of interest: time to event (T^*)
- T^* is **left truncated** by the enrollment time (Q^*) if only subjects with $T^* > Q^*$ are included in the data. \Rightarrow **Selection bias**.
 - Usually present in studies with *delayed entry*.
 - E.g., prevalent cohort studies with follow-up, aging studies, pregnancy studies.
- A toy example of aging study:

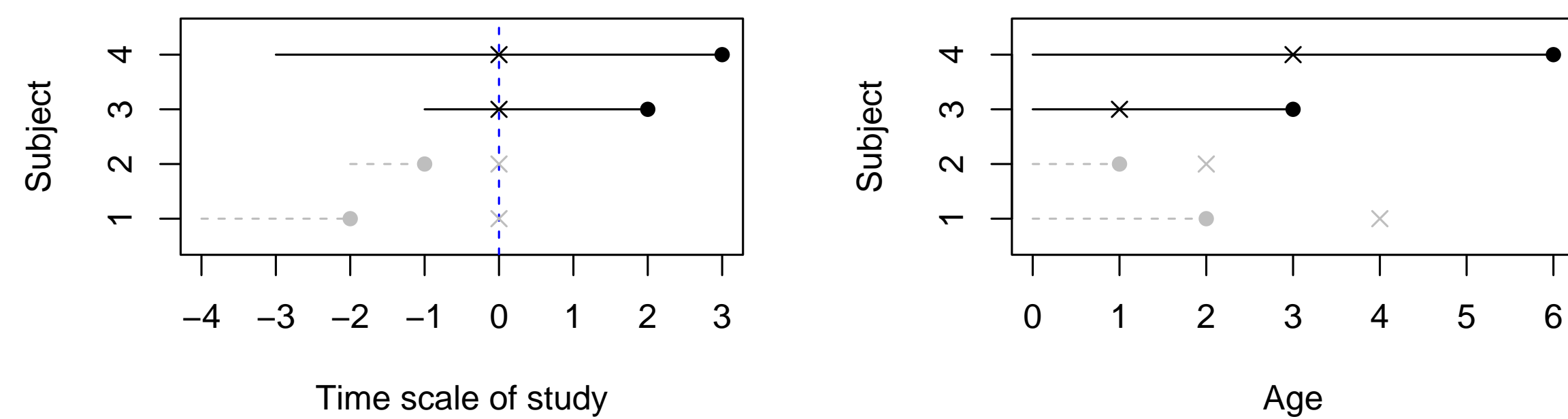
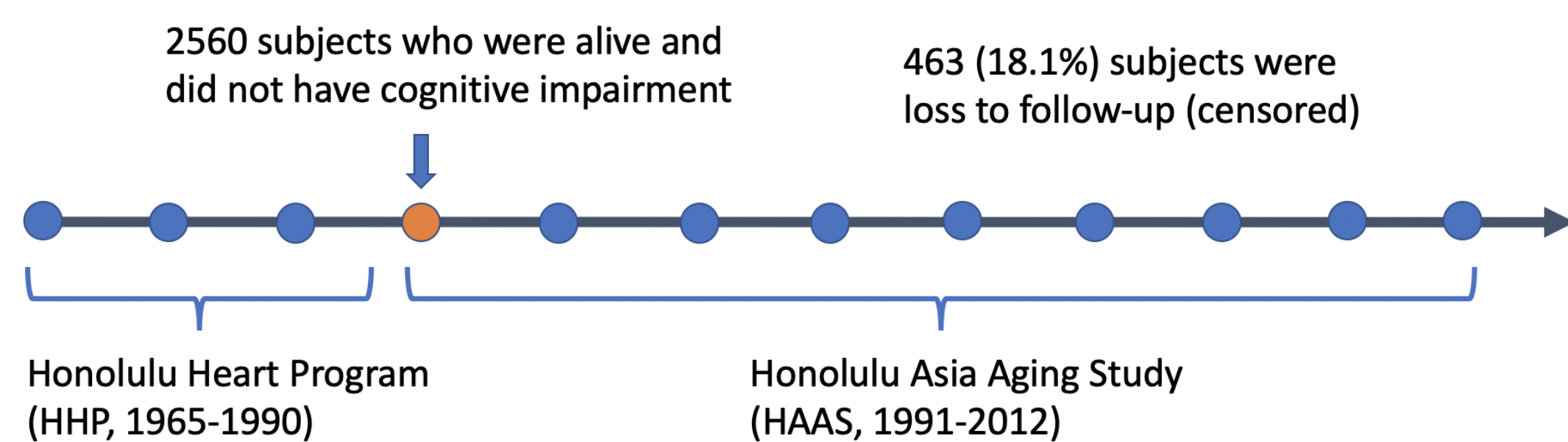


Figure: People's lifespans on the time scale of study (left) and on the age scale (right). Solid dots: event times; 'x': enrollment times.

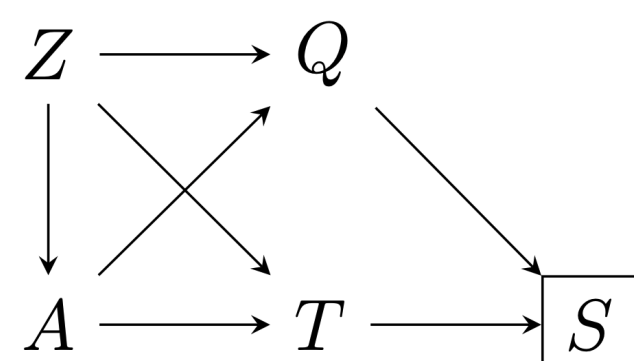
- HAAS data:



Outcome of interest: Cognitive impairment-free survival on the age scale. T^* - age to moderate cognitive impairment or death. Q^* - age at entry of HAAS.

- **Triple biases:**

- Confounding in observational data.
- **Selection bias from left truncation**
 - Early event times are underrepresented.
- Informative right censoring.



EIC for $\mathbb{E}\{\nu(T^*)\}$ under left truncation

- The efficient influence curve (EIC) for $\theta = \mathbb{E}\{\nu(T^*)\}$: $\beta \cdot U(\theta; F, G)$,

$$U(\theta; F, G) = \frac{\nu(T) - \theta}{G(T|A, Z)} - \int_0^\infty m_\nu(v, A, Z; \theta, F) \cdot \frac{F(v|A, Z)}{1 - F(v|A, Z)} \cdot \frac{d\bar{M}_Q(v; G)}{G(v|A, Z)},$$
 where $\beta = \mathbb{P}(Q^* < T^*)$, $G(t|a, z) = \mathbb{P}(Q^* \leq t | A^* = a, Z^* = z)$, $F(t|a, z) = \mathbb{P}(T^* \leq t | A^* = a, Z^* = z)$, and $m_\nu(v, a, z; \theta, F) = \mathbb{E}\{\nu(T^*) - \theta | T^* < v, A^* = a, Z^* = z\}$.
- **Double robustness:** $\mathbb{E}\{U(\theta_0; F, G)\} = 0$ if either $F = F_0$ or $G = G_0$, where F_0, G_0 denote the truth.
- Constructed *model doubly robust* estimators under asymptotic linearity and *rate doubly robust* estimators under cross-fitting.
- Provided technical conditions for the asymptotic properties that appear to not have been carefully examined in the literature for time-to-event data.
- Represents the **first attempt** to construct doubly robust estimators in the presence of left truncation (selection bias)
 - \rightarrow Does *NOT* fall under the established framework of coarsened data where doubly robust approaches are developed.

AIPTW and AIPCW operators

For any $\zeta_1(T^*(a), A^*, Z^*)$ and any $\zeta_2(Q, T, A, Z)$,

$$\mathcal{V}_a(\zeta_1; \pi, \mu_a) = \frac{(A^*)^a (1 - A^*)^{1-a}}{\pi(Z^*)^a \{1 - \pi(Z^*)\}^{1-a}} \zeta_1(T^*, A^*, Z^*) + \frac{(-1)^a \{A^* - \pi(Z^*)\}}{\tilde{\pi}(a, Z^*)} \mathbb{E}\{\zeta_1(T^*, A^*, Z^*) | A^* = a, Z^*\}$$

$$\mathcal{V}_C(\zeta_2; F, S_D)(Q, X, \Delta, A, Z) = \frac{\Delta \zeta(Q, X, A, Z)}{S_D(X - Q | Q, A, Z)} + \int_0^\infty \mathbb{E}\{\zeta_\oplus(Q, T, A, Z) | T - Q \geq u, Q, A, Z\} \frac{dM_D(u; S_D)}{S_D(u | Q, A, Z)}$$

truncAIPW operator

- For any function $\zeta(T^*, Z^*; \theta)$ of the truncation free data, consider the following truncation AIPW operator:

$$\mathcal{V}_Q(\zeta; F, G)(Q, T, A, Z) = \frac{\zeta(T, A, Z; \theta)}{G(T|A, Z)} - \int_0^\infty m_\zeta(v, A, Z; \theta, F) \cdot \frac{F(v|A, Z)}{1 - F(v|A, Z)} \cdot \frac{d\bar{M}_Q(v; G)}{G(v|A, Z)},$$

where $m_\zeta(v, a, z) = \mathbb{E}\{\zeta(T^*, Z^*; \theta) | T^* < v, A^* = a, Z^* = z\}$.

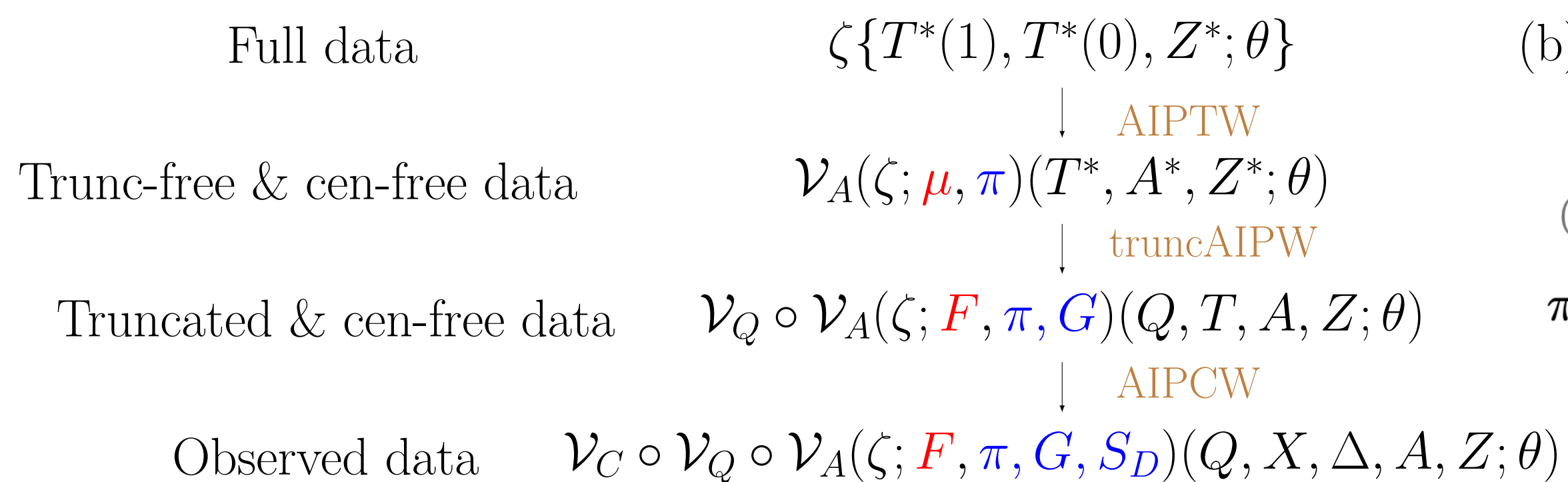
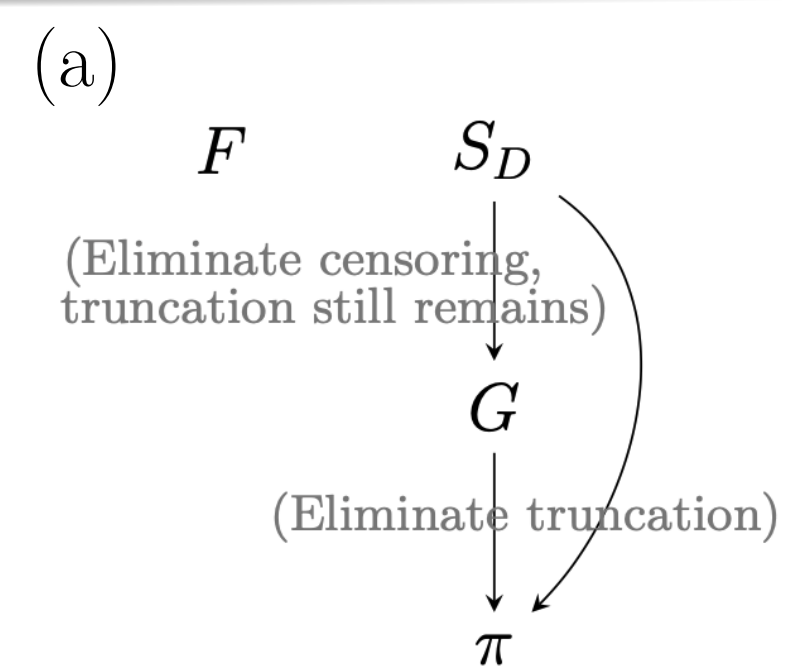
- **Double robustness:** if either $F = F_0$ or $G = G_0$, then

$$\mathbb{E}\{\mathcal{V}_Q(\zeta; F, G)(Q, T, A, Z)\} = \beta^{-1} \cdot \mathbb{E}\{\zeta(T^*, A^*, Z^*; \theta)\},$$

where $\beta = \mathbb{P}(Q^* < T^*)$.

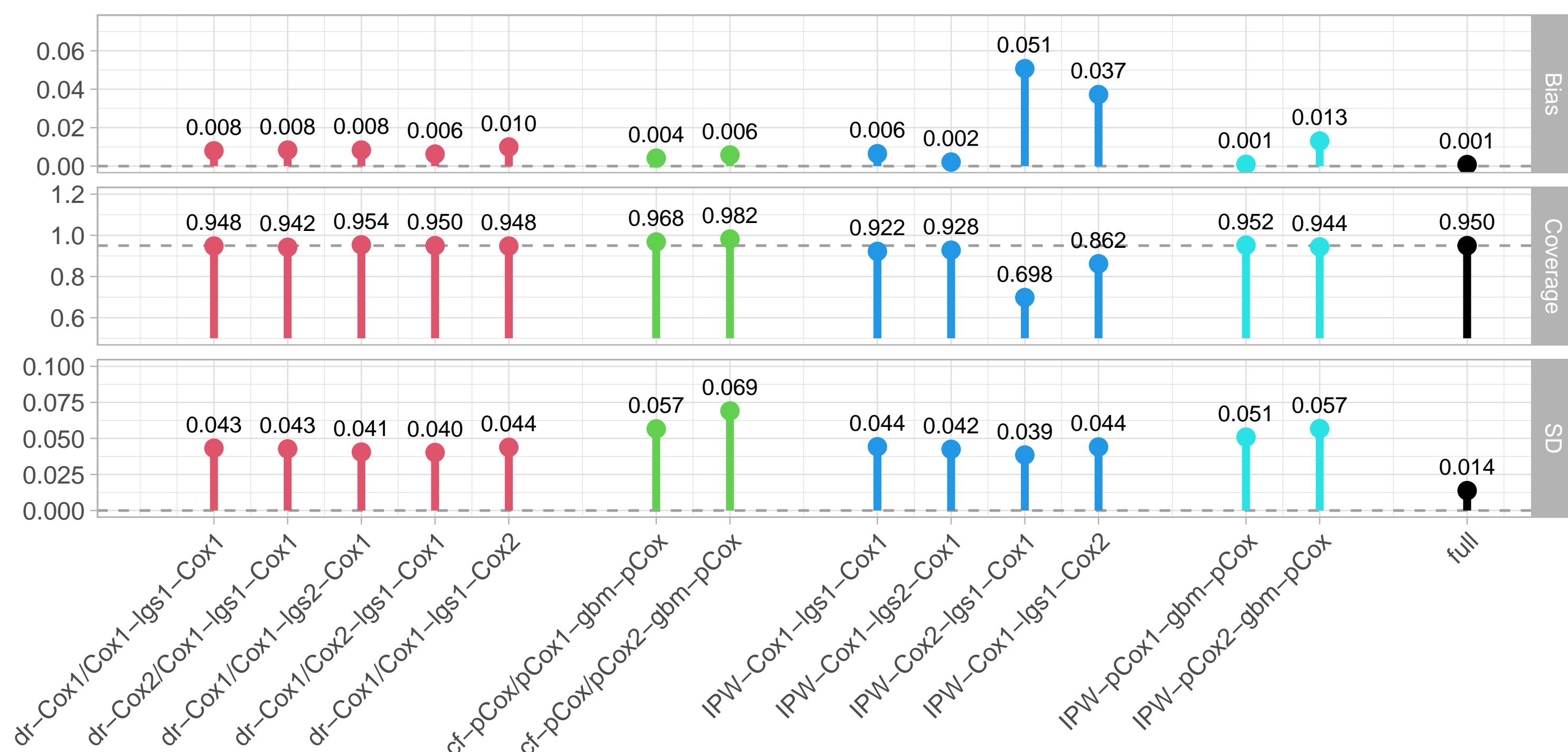
Framework for handling triple biases

- Estimand θ - defined by the distribution of $(T^*(1), T^*(0), Z^*)$; e.g. ATE: $\mathbb{E}\{\nu\{T^*(1)\} - \nu\{T^*(0)\}\}$.
- $A \in \{1, 0\}$ - treatment. $T^* = A^*T^*(1) + (1 - A^*)T^*(0)$.
- D - residual censoring. $X = \min(T, Q + D)$, $\Delta = I(T < Q + D)$.



Simulation results

- 500 simulated data sets each with sample size 1000; estimand: $\theta = \mathbb{P}\{T^*(1) > 3\} = 0.5529$.
- Truncation rate: 22.8%; treatment rate: 50%; censoring rate 47.6%.



Orthogonal learning for HTE

- Conditional average treatment effect (CATE): $\tau_0(z) = \mathbb{E}\{\nu\{T^*(1)\} - \nu\{T^*(0)\} | Z^* = z\}$
- Loss function for truncation-free and censoring-free data: $\ell(T^*, A^*, Z^*; \tau, m, \pi)$; e.g., R-loss (Nie and Wager 2021), DR-loss (Kennedy 2023), etc.
- Apply $\mathcal{V}_C \circ \mathcal{V}_Q$ to ℓ :

$$L(\tau; m, \pi, F, G, S_D) = \mathcal{V}_C \circ \mathcal{V}_Q(\ell; F, \pi, G, S_D)(Q, X, \Delta, A, Z).$$

We have

$$\tau_0 = \arg \min_{\tau} \mathbb{E}\{L(\tau; F, \pi, G, S_D)\}.$$

- L is a Neyman orthogonal loss
 - \Rightarrow Achieve oracle rate if the nuisance parameters are estimated at faster than $n^{-1/4}$ rate.

Additional Information

- Wang, Y., Ying, A., & Xu, R. (2024). Doubly robust estimation under covariate-induced dependent left truncation. *Accepted in Biometrika*. arXiv:2208.06836.
- Code: https://github.com/wangyuyao98/left_trunc_DR.
- Email: yuw079@ucsd.edu (Yuyao Wang).